## Exercise H1.1

a) Suppose the data $X$ in a statistical model take values in a countable set $\mathcal{X}$ (i.e. $X$ has a discrete law). In class it was claimed that the data itself are a sufficient statistic (i.e. $T(X)=X$ is sufficient). Write down the argument (it can be a very short paragraph).
b). Suppose $T$ is a sufficient statistic with values in a set $\mathcal{T}$ and $S: \mathcal{T} \mapsto \mathcal{S}$ is a mapping with values in a set $\mathcal{S}$ which is one-to-one (i.e there exists an inverse mapping $S^{-1}$ such that $S^{-1}(S(t))=t$ for all $\left.t \in \mathcal{T}\right)$ Show that the statistic $S(T(X))$.is sufficient.
c) In example 2.1 handout it was claimed that the statistic $T(X)=\left(X_{1}, \bar{X}_{n}\right)$ is sufficient in Model I. Prove this claim.

Exercise H1.2.Let $X_{1}, \ldots, X_{n}$ be independent and identically distributed with Poisson law $\operatorname{Po}(\lambda)$, where $\lambda>0$ is unknown. Show that the sample mean $\bar{X}_{n}$ is again a sufficient statistic (Comment: the sample mean is a sufficient statistic not only for i.i.d. Bernoulli data, but in a number of statistical models).

Exercise H1.3. Let $X_{1}, \ldots, X_{n}$ be independent and identically distributed such that $X_{1}$ has the uniform law on the set $\{1, \ldots, r\}$ for some integer $r>1$ (i.e. $P\left(X_{1}=k\right)=1 / r, k=$ $1, \ldots, r)$. In the statistical model where $r>1$ is unknown, show that $T(X)=\max _{i=1, \ldots, n} X_{i}$ is a sufficient statistic. Hint: as an intermediate step, show that $P(T(X) \leq k)=(k / r)^{n}$ for all $k=1, \ldots, r$. (Comment: in this model, the unknown parameter is the largest value that the data can possibly take, i.e. $r$. It turns out that the largest value which they take in the sample is a sufficient statistic.).

Exercise H1.4. Let $X_{1}, \ldots, X_{n}$ be independent and identically distributed such that $X_{1}$ has the geometric law $\operatorname{Geom}(p)$, i.e.

$$
P\left(X_{1}=k\right)=(1-p)^{k-1} p, k=1,2, \ldots
$$

for some $p \in(0,1)$. In the statistical model where $p$ is unknown, show that the sample mean $\bar{X}_{n}$ is a sufficient statistic. Hint: it can be used that $n \bar{X}_{n}=\sum_{i=1}^{n} X_{i}$ has the negative binomial distribution with parameters $n$ and $p$, i.e.

$$
P\left(n \bar{X}_{n}=k\right)=\binom{k-1}{k-n}(1-p)^{k-n} p^{n} \text { for } k \geq n
$$

## Exercise H1.1.

a) $P_{\theta}(X \in B \mid X=x)=\frac{P_{\theta}(X \in B, X=x)}{P_{\theta}(X=x)}=1_{B}(x)$, independent of $\theta$.
b) $P_{\theta}(X \in B \mid T(X)=t)$ is independent of $\theta$ for any event $B$ and any value $t$ of the set $\mathcal{T}$, since $T$ is a sufficient statistic. This implies $P_{\theta}(X \in B \mid S(T(X))=s)=$ $P_{\theta}\left(X \in B \mid T(X)=S^{-1}(s)\right)$ is independent of $\theta$ for any event $B$ and any value $s$ of the set $\mathcal{S}$. Because $S: \mathcal{T} \longrightarrow \mathcal{S}$ is one to one and onto, $S^{-1}(s) \in \mathcal{T}$. for any $s \in \mathcal{S}$.
c) We will show that $P_{p}\left(X_{1}=x_{1}, X_{2}=x_{2}, \ldots, X_{n}=x_{n} \mid X_{1}=x_{1}^{\prime}, n \overline{X_{n}}=k\right)$ is independent of $p$ for any $\left(x_{1}, \ldots, x_{n}\right) \in\{0,1\}^{n}, x_{1}^{\prime} \in\{0,1\}$, and $0 \leqslant k \leqslant n$. If $x_{1} \neq x_{1}^{\prime}$, or $k \neq n \overline{x_{n}}$, the conditional probability is 0 . If $x_{1}=x_{1}^{\prime}$, and $k=n \overline{x_{n}}$, then
$P_{p}\left(X_{1}=x_{1}^{\prime}, n \overline{X_{n}}=k\right)=P_{p}\left(X_{1}=x_{1}, X_{2}+\ldots+X_{n}=x_{2}+\ldots+x_{n}\right)$ $=P_{p}\left(X_{1}=x_{1}\right) P_{p}\left(X_{2}+\ldots+X_{n}=x_{2}+\ldots+x_{n}\right)$

$p)^{(n-1)-\left(x_{2}+\ldots+x_{n}\right)}$

$$
=\binom{n-1}{x_{2}+\ldots+x_{n}} p^{\sum_{i=1}^{n} x_{i}}(1-p)^{n-\sum_{i=1}^{n} x_{i}}
$$

and,

$$
\begin{aligned}
P_{p}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}, X_{1}=x_{1}, n \overline{X_{n}}=k\right)= & P_{p}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}\right) \\
& =\underline{p^{\sum_{i=1}^{n} x_{i}}(1-p)^{n-\sum_{i=1}^{n} x_{i}}}
\end{aligned}
$$

This implies $P_{p}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid X_{1}=x_{1}, n \overline{X_{n}}=n \overline{x_{n}}\right)=1 /\binom{n-1}{x_{2}+\ldots+x_{n}}$ is independent of $p$.

So $\left(X_{1}, n \overline{X_{n}}\right)$ is a sufficient stattistic, i.e., $\left(X_{1}, \overline{X_{n}}\right)$ is sufficient from (b).

## Exercise H1.2.

We need to prove that $\overline{X_{n}}$ is sufficcient, i.e., $n \overline{X_{n}}$ is sufficient from(b), i.e., $P_{\lambda}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid n \overline{X_{n}}=k\right)$ is independent of $\lambda$ for any $\left(X_{1}, \ldots, X_{n}\right) \in$ $\{0,1,2, \ldots\}^{n}$, and $k \geqslant 0$. If $k \neq \sum_{i=1}^{n} x_{i}$, the conditional probability is 0 . If $k=\sum_{i=1}^{n} x$, we know that the distribution of $\sum_{i=1}^{n} x_{i}$ is $\operatorname{Po}(n \lambda)$,
then

$$
P_{\lambda}\left(n \overline{X_{n}}=n \overline{x_{n}}\right)=e^{-n \lambda \frac{(n \lambda)^{\sum_{i=1}^{n} x_{i}}}{\left(\sum_{i=1}^{n} x_{i}\right)!},}
$$

and,

$$
\begin{aligned}
P_{\lambda}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}, n \overline{X_{n}}=k\right)=P_{p} & \left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}\right) \\
= & \prod_{i=1}^{n}\left(e^{-\lambda} \frac{\lambda^{x_{i}}}{x_{i}!}\right) \\
& =e^{-n \lambda \frac{\lambda^{\sum_{i=1}^{n} x_{i}}}{\prod_{i=1}^{n} x_{i}!}}
\end{aligned}
$$

This implies $P_{\lambda}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid n \overline{X_{n}}=n \overline{x_{n}}\right)=\left(\sum_{i=1}^{n} x_{i}\right)!/\left(n^{\sum_{i=1}^{n} x_{i}} \prod_{i=1}^{n}\left(x_{i}!\right)\right)$ is independent of $\lambda$. So $\overline{X_{n}}$ is sufficient.

Exercise H1.3.

We need to show that $P_{r}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid T(x)=\max _{i=1, \ldots n} X_{i}=k\right)$ is independent of $r$ for any $x_{1}, \ldots, x_{n}$, and $k$. If $\max _{i=1, \ldots, n} x_{i} \ngtr k$, the conditional probability is 0 . If $\max _{i=1, \ldots, n} x_{i} \leqslant k$, we have
$P_{r}\left(T(x)=\max _{i=1, \ldots, n} x_{i}=k\right)=P_{r}(T(x) \leqslant k)-P_{r}(T(x) \leqslant k-1)$

$$
=P_{r}\left(X_{1} \leqslant k, \ldots, X_{n} \leqslant k\right)-
$$

$P_{r}\left(X_{1} \leqslant k-1, \ldots, X_{n} \leqslant k-1\right)$

$$
\begin{gathered}
=\prod_{i=1, \ldots, n}^{n} P_{r}\left(X_{i} \leqslant k\right)-\prod_{i=1, \ldots, n}^{n} P_{r}\left(X_{i} \leqslant k-1\right) \\
=\left(\frac{k}{r}\right)^{n}-\left(\frac{k-1}{r}\right)^{n},
\end{gathered}
$$

and,

$$
\begin{gathered}
P_{r}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}\right)=\prod_{i=1, \ldots, n}^{n} P_{r}\left(X_{i}=x_{i}\right) \\
=\left(\frac{1}{r}\right)^{n} .
\end{gathered}
$$

This implies $P_{r}=\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid T(x)=\max _{i=1, \ldots, n} x_{i}\right)=\frac{\left(\frac{1}{r}\right)^{n}}{\left(\frac{k}{r}\right)^{n}-\left(\frac{k-1}{r}\right)^{n}}=$ $\frac{1}{k^{n}-(k-1)^{n}}$ is independent of $r$.

So $T(x)$ is sufficient.

Exercise H1.4.
We need to show $\overline{X_{n}}$ is sufficient, i.e., $n \overline{X_{n}}$ is sufficient, i.e., $P_{p}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid n \overline{X_{n}}=k\right)$ is independent of $p$ for any $\left(x_{1}, \ldots, x_{n}\right) \in\{1,2,3 \ldots .\}^{n}$, and $k \geqslant n$. If $k \neq n \overline{x_{n}}$, then the conditinal probability is 0 . If $k=n \overline{x_{n}}$, then

$$
P_{p}\left(n \overline{X_{n}}=n \overline{x_{n}}\right)=\left(\frac{n \overline{x_{n}}-1}{n \overline{x_{n}}-n}\right)(1-p)^{n \overline{x_{n}}-n} p^{n}
$$

## and,

$$
\begin{aligned}
P_{p}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}, n \overline{X_{n}}=n \overline{x_{n}}\right)=P_{p} & \left(X_{1}=x_{1}, \ldots, X_{n}=x_{n}\right) \\
& =\prod_{i=1}^{n}(1-p)^{x_{i}-1} p \\
& =(1-p)^{n \overline{x_{n}}-n} p^{n} .
\end{aligned}
$$

This implies $P_{p}\left(X_{1}=x_{1}, \ldots, X_{n}=x_{n} \mid n \overline{X_{n}}=n \overline{x_{n}}\right)=1 /\left(\frac{n \overline{x_{n}}-1}{n \overline{x_{n}}-n}\right)$ is independent of $r$.

So $\overline{X_{n}}$ is sufficient.

