# Multivariate Statistical Analysis Fall 2011

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### Lecture 9 for Applied Multivariate Analysis

# Outline



- T<sup>2</sup> distribution in the two sample case
- Wilk's Lambda



Analogous to the univariate context, we wish to determine whether the mean vectors are comparable, more formally:

$$H_0: \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \tag{1}$$

Suppose we let  $\mathbf{y}_{1i}$ ,  $i = 1, ..., n_1$  and  $\mathbf{y}_{2i}$ ,  $i = 1, ..., n_2$  represent independent samples from two *p*-variate normal distribution with mean vectors  $\boldsymbol{\mu}_1$  and  $\boldsymbol{\mu}_2$  but with common covariance matrix  $\boldsymbol{\Sigma}$ unknown, provided  $\boldsymbol{\Sigma}$  is positive definite and n > p, given sample estimators for mean and covariance  $\bar{\mathbf{y}}$  and  $\mathbf{S}$  respectively. We can then define

$$\mathbf{W}_{1} = (n_{1} - 1) \mathbf{S}_{1} = \sum_{i=1}^{n_{1}} (\mathbf{y}_{1i} - \bar{\mathbf{y}_{1}}) (\mathbf{y}_{1i} - \bar{\mathbf{y}_{1}})'$$
$$\mathbf{W}_{2} = (n_{2} - 1) \mathbf{S}_{2} = \sum_{i=1}^{n_{2}} (\mathbf{y}_{2i} - \bar{\mathbf{y}_{2}}) (\mathbf{y}_{2i} - \bar{\mathbf{y}_{2}})'$$

since each are unbiased estimators of the common covariance matrix, ie.  $E[(n_1 - 1) \mathbf{S}_1] = (n_1 - 1) \mathbf{\Sigma}$  and  $E[(n_2 - 1) \mathbf{S}_2] = (n_2 - 1) \mathbf{\Sigma}$ 

The  $T^2$  statistic can be calculated as:

$$T^{2} = \left(\frac{n_{1}n_{2}}{n_{1}+n_{2}}\right) \left(\frac{n_{1}n_{2}}{n_{1}+n_{2}}\right) \left(\mathbf{\bar{y}}_{1}-\mathbf{\bar{y}}_{2}\right)' \mathbf{S}^{-1} \left(\mathbf{\bar{y}}_{1}-\mathbf{\bar{y}}_{2}\right)$$
(2)

where  $\mathbf{S}^{-1}$  is the inverse of the pooled correlation matrix given by:

$$S = \frac{(n_1 - 1)S_1 + (n_2 - 1)S_2}{n_1 + n_2 - 2}$$
$$= \frac{1}{n_1 + n_2 - 2} (W_1 + W_2)$$

given the sample estimates for covariance,  $\boldsymbol{S}_1$  and  $\boldsymbol{S}_2$  in the two samples.

 $\mathsf{T}^2$  distribution in the two sample case Wilk's Lambda

# Outline

# Two sample T<sup>2</sup> test T<sup>2</sup> distribution in the two sample case Wilk's Lambda

2 Confidence ellipses

 Two sample T<sup>2</sup> test
 T<sup>2</sup> distribution in the two sample case

 Confidence ellipses
 Wilk's Lambda

Again, there is a simple relationship between the test statistic,  $T^2$ , and the F distribution:

#### Theorem

If  $\mathbf{y}_{1i}$ ,  $i = 1, ..., n_1$  and  $\mathbf{y}_{2i}$ ,  $i = 1, ..., n_2$  represent independent samples from two p variate normal distribution with mean vectors  $\mu_1$  and  $\mu_2$  but with common covariance matrix  $\boldsymbol{\Sigma}$ , provided  $\boldsymbol{\Sigma}$  is positive definite and n > p, given sample estimators for mean and covariance  $\bar{\mathbf{y}}$  and  $\mathbf{S}$  respectively, then:

$$F = \frac{(n_1 + n_2 - p - 1)T^2}{(n_1 + n_2 - 2)p}$$

has an F distribution on p and  $(n_1 + n_2 - p - 1)$  degrees of freedom.

- Essentially, we compute the test statistic, and see whether it falls within the  $(1 \alpha)$  quantile of the F distribution on those degrees of freedom.
- Note again that to ensure non-singularity of **S**, we require that  $n_1 + n_2 > p$ .

 $\mathsf{T}^2$  distribution in the two sample case Wilk's Lambda

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## Characteristic form

$$T^{2} = (\mathbf{\bar{y}}_{1} - \mathbf{\bar{y}}_{2})' \left[ \left( \frac{1}{n_{1}} \frac{1}{n_{2}} \right) \mathbf{S}_{\rho l} \right]^{-1} (\mathbf{\bar{y}}_{1} - \mathbf{\bar{y}}_{2})$$
(3)

 $\mathsf{T}^2$  distribution in the two sample case Wilk's Lambda

# Outline



•  $T^2$  distribution in the two sample case

Wilk's Lambda

2 Confidence ellipses

What was all that stuff about likelihood ratio's about? It turns out that it is possible to show that:

$$\Lambda^{2/n} = \left(\frac{|\hat{\boldsymbol{\Sigma}}|}{|\hat{\boldsymbol{\Sigma}}_0|}\right) = \left(1 + \frac{T^2}{n-1}\right)^{-1}$$
(4)

It is also possible to obtain the  $T^2$  via union intersection methods. This is nice because it tells us a lot about the properties of the test!

# Confidence ellipses

Essentially, we wish to find a region of squared Mahalanobis distance such that:

$$\mathsf{Pr}\left((\mathbf{ar{y}}-oldsymbol{\mu})'\mathbf{S}^{-1}(\mathbf{ar{y}}-oldsymbol{\mu})
ight)\leq c^2$$

and we can find  $c^2$  as follows:

$$c^{2} = \left(\frac{n-1}{n}\right) \left(\frac{p}{n-p}\right) F_{(1-\alpha),p,(n-p)}$$

where  $F_{(1-\alpha),p,(n-p)}$  is the  $(1-\alpha)$  quantile of the *F* distribution with *p* and *n* - *p* degrees of freedom, *p* represents the number of variables and *n* the sample size.



- $\bullet\,$  The centroid of the ellipse is at  $\bar{y}$
- The half length of the semi-major axis is given by:

$$\sqrt{\lambda_1}\sqrt{\frac{p(n-1)}{n(n-p)}}F_{p,n-p}(\alpha)$$

where  $\lambda_1$  is the first eigenvalue of  $\boldsymbol{S}$ 

• The half length of the semi-minor axis is given by:

$$\sqrt{\lambda_2}\sqrt{\frac{p(n-1)}{n(n-p)}}F_{p,n-p}(\alpha)$$

where  $\lambda_2$  is the second eigenvalue of  $\boldsymbol{S}$ 

• The ratio of these two eigenvalues gives you some idea of the elongation of the ellipse



- In addition to the (joint) confidence ellipse, it is possible to consider *simultaneous* confidence intervals - univariate confidence intervals based on a linear combination which could be considered as shadows of the confidence ellipse
- It is also possible to carry out Bonferroni adjustments of these simultaneous intervals



- T<sup>2</sup> test is based upon Mahalanobis distance and can be used for inference on mean vectors - this test can be derived via a variety of routes
- Difference between univariate and multivariate inference, especially when considering confidence ellipses
- Having determined that there is a significant difference between mean vectors, you may wish to conduct a number of follow up investigations and even carry out discriminant analysis